

# Nonlinear structural equation modeling

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# Structure

- 1 Introduction
- 2 Product indicator approaches
  - Constrained approach
  - Means and (co)variances of product variables
  - Unconstrained pi approach
- 3 Alternative approaches for nonlinear SEM
  - LMS
  - QML
  - 2SMM
  - Application of LMS, QML, and 2SMM
- 4 Summary
- 5 Some simulation results for different approaches

# Introduction

# Linear models: Limitations

- Linear models such as regression models or SEM often assume linear relationships between the variables.
- This assumption is often either wrong or a very crude approximation of the actual functional relationship.
- From an applied perspective, interaction effects are important because they indicate if the relationship between two variables is moderated by a third variable.
- In this workshop, we will discuss a variety of options to analyze deviations from linearity in the latent variable framework.

# Why latent variables?

- Latent variable models, or more specifically here: Structural equation models (SEM) allow to consider a variety of aspects in data analysis:
  - They take measurement error into account by using multiple indicators and by specifying relationships between latent variables that are theoretically measurement error free
  - More complex relationships can be analyzed
  - Model fit can be investigated
- Particularly for nonlinear effects, the measurement error aspect is important because a typical effect size in social sciences is around 2.5%.

# What is nonlinear?

- Nonlinearity can involve very different aspects:
  - Interaction effects between latent variables
  - Curvilinear relationships between latent variables
  - Nonlinear measurement models, for example for dichotomous or count data
  - Nonlinear parameters (such as  $\lambda^2$ )
- In this workshop, we focus on the first two aspects.

# Interaction effects

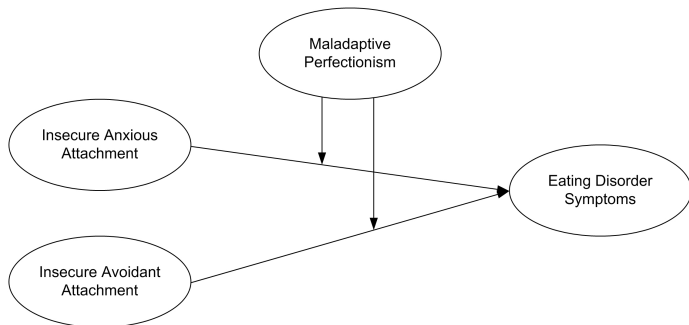
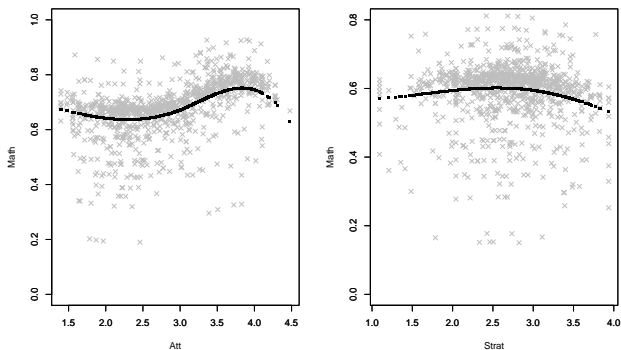


Figure: Taken from Dakanalis et al. (2014)

# Quadratic and other polynomial effects



**Figure:** Taken from Kelava and Brandt (2015). Relationship between between pupils' math skills (Math) and their attitude toward reading (Att; left), and the reported teaching strategies (Strat; right).



# Overview workshop

- 1 Interaction effects: Traditional product indicator approaches
- 2 Interaction and quadratic effects: Other approaches
- 3 Mixture models for curvilinear relationships
- 4 Mixture models for interaction and quadratic effects
- 5 Multilevel models with interaction effects

# Overview sem-3

- 1 Interaction effects: Traditional product indicator (PI) approaches
  - Constrained approach
  - GAPI approach
  - Unconstrained approach
  - Main package: lavaan
- 2 Interaction and quadratic effects: Other approaches
  - Moment-based approaches
  - “Distribution-analytic” maximum likelihood approaches (LMS)
  - Standardization and illustration of results
  - Main package: nlsem and experimental syntax for 2SMM

# Overview sem-4

- 1 Structural equation mixture models (SEMM)
  - Direct and indirect applications
  - Growth curve mixture models (direct application)
  - Optional: Alternative models for heterogeneous growth
  - Example for curvilinear relationship (indirect application)
- 2 Background: EM algorithm
- 3 Nonlinear structural equation mixture models (NSEMM)
  - Standardization of effects in NSEMM
- 4 Main software: Mplus, nlsem and plotsemm

# Overview sem-5

- 1 Multilevel models with interaction effects
  - Bayesian modeling (priors, logic of Bayesian modeling)
  - Multilevel models with and without interaction effects
  - Multilevel SEM with and without interaction effects
  - Main packages: rstan and stan

# Regression model

- In regression models, the inclusion of an interaction effect is typically conducted by including a product term of two (or more) predictor variables:

$$y_i = \beta_0 + \beta_1 x_{1i} + \beta_2 x_{2i} + \beta_3 x_{1i} x_{2i} + \epsilon_i \quad (1)$$

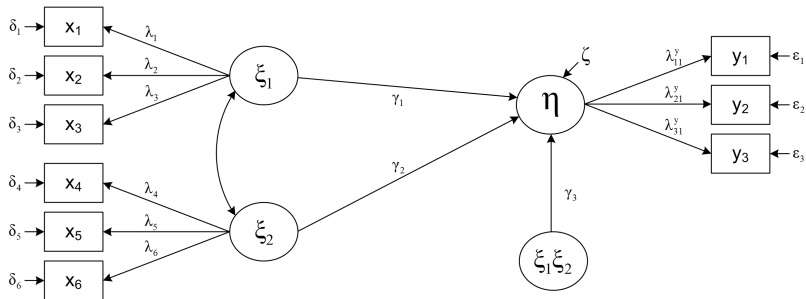
- A reformulation shows the moderation:

$$y_i = \beta_0 + \beta_1 x_{1i} + \underbrace{(\beta_2 + \beta_3 x_{1i})}_{\beta_{2i}^*} x_{2i} + \epsilon_i \quad (2)$$

The relationship between  $x_2$  and  $y$  depends on the value  $x_{1i}$  that may be different for each person  $i$ .

- The model is symmetric, i.e. the decision if  $x_1$  or  $x_2$  is the moderator is not a statistical but only applied question.

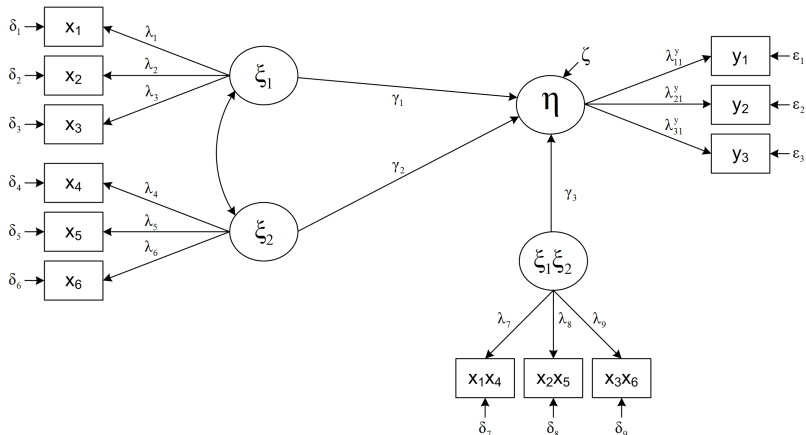
## SEM with interaction effect



## Problem of SEM: product terms

- The main question is: how can one form the product of  $\xi_1$  and  $\xi_2$ ?
- Intuitive solution (Kenny & Judd, 1984): build products of indicators for  $\xi_1$  and  $\xi_2$  and use them as “product indicators” for a latent interaction term  $\xi_1 \cdot \xi_2$ .
- Use standard SEM software to estimate the model (e.g., LISREL)

## SEM with interaction effect and pi's





# The constrained pi approach (Kenny & Judd, 1984)

- The first approach used the following logic: The parameters from the measurement model for the pi's are functions of the parameters of measurement model of the indicator. By constraining them accordingly, the model is still parsimonious.
- How does one derive these constraints?

$$x_1 \cdot x_4 = \lambda_7 \xi_1 \xi_2 + \delta_7 \quad (3)$$

- Example:

$$x_1 = \lambda_1 \xi_1 + \delta_1 \quad (4)$$

$$x_4 = \lambda_4 \xi_2 + \delta_4 \quad (5)$$

- And so

$$x_1 \cdot x_4 = (\lambda_1 \xi_1 + \delta_1) \cdot (\lambda_4 \xi_2 + \delta_4) \quad (6)$$

# The constrained pi approach (Kenny & Judd, 1984)

- Based on this equation one can derive

$$x_1 \cdot x_4 = (\lambda_1 \xi_1 + \delta_1) \cdot (\lambda_4 \xi_2 + \delta_4) \quad (7)$$

$$= \underbrace{\lambda_1 \lambda_4}_{=\lambda_7} \xi_1 \xi_2 + \underbrace{\lambda_1 \xi_1 \delta_4 + \lambda_4 \xi_2 \delta_1 + \delta_1 \delta_4}_{=\delta_7} \quad (8)$$

- Factor loading constraint:

$$\lambda_7 = \lambda_1 \lambda_4 \quad (9)$$

# The constrained pi approach (Kenny & Judd, 1984)

- Residual variance constraint

$$Var(\delta_7) = Var(\lambda_1 \xi_1 \delta_4 + \lambda_4 \xi_2 \delta_1 + \delta_1 \delta_4) \quad (10)$$

$$= \lambda_1^2 \phi_{11} \theta_{44} + \lambda_4^2 \phi_{22} \theta_{11} + \theta_{11} \theta_{44} \quad (11)$$

- Constraints for latent variables

$$Var(\xi_1 \xi_2) = \phi_{11} \phi_{22} + \phi_{12}^2 \quad (12)$$

$$E(\xi_1 \xi_2) = \phi_{12} \quad (13)$$

$$Cov(\xi_1 \xi_2, \xi_1) = Cov(\xi_1 \xi_2, \xi_2) = 0 \quad (14)$$

- under the assumptions of
  - All latent variables have a mean of zero
  - All latent variables are normally distributed

# The mean and variance of a product variable

- Mean of the product  $X$  and  $Y$  (Bohrnstedt & Goldberger, 1969)

$$E[XY] = E[X]E[Y] + C(X, Y) \quad (15)$$

which is only zero if  $X$  and  $Y$  are uncorrelated.

- Variance is

$$\begin{aligned} Var(XY) = & E^2[X]Var(Y) + E^2[Y]Var(X) + 2E[X]E[Y]C(X, Y) \\ & + Var(X)Var(Y) + C^2(X, Y) \end{aligned} \quad (16)$$

which simplifies for centered variables to

$$Var(XY) = Var(X)Var(Y) + C^2(X, Y) \quad (17)$$

## The covariances of product variables

- The covariance between a variable  $Z$  and a product variable  $XZ$  depends on the multivariate skewness  $\nu_{xyz}$  is

$$C(XY, Z) = E[X]C(Y, Z) + E[Y]C(X, Z) + \nu_{xyz} \quad (18)$$

which is only zero if  $X$  and  $Y$  have zero means *and* data is normally distributed.

- Similar results are obtained for the covariance between two product  $XY$  and  $UV$  which has a simplified formula for *centered* variables of

$$C(XY, UV) = -C(X, Y)C(U, V) + \nu_{xyuv} \quad (19)$$

with fourth central moment (kurtosis)  $\nu_{xyuv}$  that is under normality

$$\nu_{xyuv} = C(X, Y)C(U, V) + C(X, U)C(Y, V) + C(X, V)C(Y, U) \quad (20)$$

# What does that imply for the constrained approach?

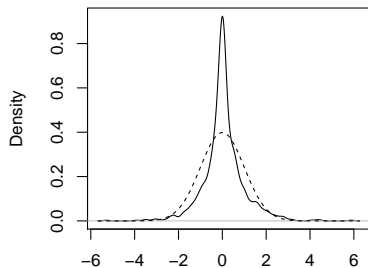
- The formulated constraints are only valid for normally distributed and centered data.
- As soon as data is nonnormal, the model is *structurally* misspecified:
  - Variances of the latent variables
  - Covariances between the latent variables

Both lead to biased estimates, also for the regression coefficients.

- But even under normality, the model leads to problems, because the normality assumption for the ML estimator is violated.

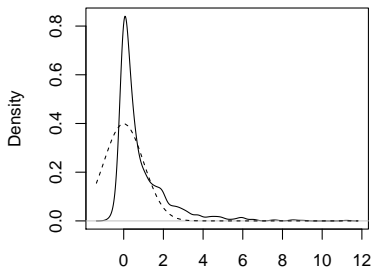
# Approach under normality

density.default(x = x[, 1] \* x[, 2])



N = 1000 Bandwidth = 0.127

density.default(x = x[, 1] \* x[, 2])



N = 1000 Bandwidth = 0.1929

**Figure:** Distribution of product variables (left: uncorrelated, right: correlated). The normality assumption (dotted lines) for the ML estimator is clearly violated.

# Demonstration 1

The constrained approach in lavaan. (for a simple model)



## Summary constrained approach

- The approach uses product indicators that form a measurement model for the latent product terms that represent the interaction effects.
- The inclusion of pi's (e.g.,  $x_1x_4$ ) lead to a violation of the normality assumption for the ML estimator which is typically used in SEM software.
- In situations with nonnormal data (e.g.,  $x_1$  is nonnormal) the approach is structurally misspecified.
- From a practical viewpoint: The formulation of constraints is extremely error prone and cumbersome.

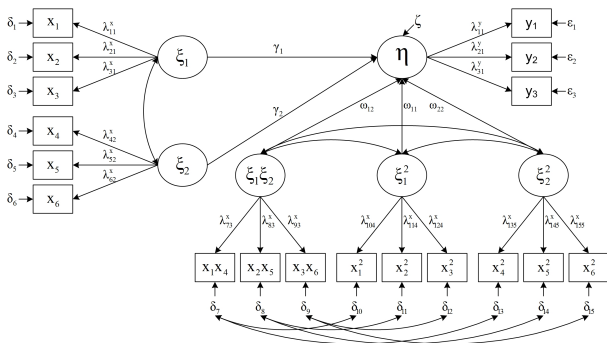
# GAPI and unconstrained pi approaches

- As a consequence to the potential problems of the constrained approach mainly two alternatives were suggested
- Wall and Amemiya (2001) proposed the Generalized Appended pi (gapi) approach that estimated the variances and covariances for the latent predictor variables freely (i.e. variances and covariances for  $\xi$  and their products)
- Marsh, Wen, and Hau (2004) proposed to estimate all parameters freely. This approach is widely known as unconstrained approach.

# Unconstrained pi approach

- The approach shows quite robust results under normal and nonnormal data conditions for small models (e.g., one interaction effect).
- Its main advantages are
  - Easy to specify
  - Can be implemented in all standard software
- On the other hand problems remain:
  - The assumptions for the ML estimator are violated. Even a robust MLR estimator with sandwich type standard errors leads to increased type I error rates.
  - If data is nonnormal (e.g.,  $x_1$ ), covariances are misspecified because they are assumed to be zero, e.g. between  $C(\xi_1, \xi_1\xi_2)$  or  $C(\delta_1, \delta_7)$
  - If models are more complex, additional error covariances need to be estimated. . .
  - Selection of product terms

# Unconstrained pi approach



**Figure:** Unconstrained PI with two quadratic and one interaction effect. Additional error covariances need to be estimated (Kelava & Brandt, 2009).

# Selection of indicators

- When forming product indicators, different possibilities exist:
  - Use non-redundant pairs indicators (only applicable for equal numbers of indicators per construct). This method is preferred.
  - Use all possible pairs. This leads to a very complex model and additional residual covariances need to be estimated freely (Kelava & Brandt, 2009)
- Even with non-redundant pairs, the model becomes fast very complex and instable when including more than two or three nonlinear effects.

## Exercise II

- 1 Estimate the interaction model for the Kenny-Judd data using the upi approach.
- 2 Compare the findings of the cpi and the upi approach.
- 3 How do results change depending on the type of standard error estimation?

## Exercise III

- The dataset was taken from PISA 2009 (OECD, 2010)
- Australian subsample of  $N = 1,019$  students who took part in a reading test.
- Constructs:
  - *Predictor 1*: Students' attitude towards reading (*Att*)
  - *Predictor 2*: Reported online activities (*Onl*; i.e., read emails or chat online)
  - *Dependent variable*: Reading skills (*Read*)
- For each latent construct 3 item parcels were constructed that are saved in the file `pisa_online.dat`.

## Exercise III

- 1 Investigate the data (distribution etc.)
- 2 Visualize the model in a path diagram. Include a model with an interaction and two quadratic effects.
- 3 Extend the R syntax used before and illustrate the results using the code available.



# Latent moderated structures (LMS)

- 1 Klein and Moosbrugger (2000) developed a maximum likelihood estimator
- 2 This approach assumes normal data for the  $x$  variables (indicators of latent predictors) only.
- 3 No need for products of indicators.
- 4 The model takes the specific nonlinear structure of the means and covariances into account for estimation
- 5 Special software necessary
  - Mplus (Muthén & Muthén, 1998–2012)
  - R package `nlssem` (Umbach, Naumann, Brandt, & Kelava, 2017)

# LMS: Background

- The main idea for LMS is to formulate a likelihood function for the observed data vector  $(x, y)$ . For this, conditional means and covariance matrices need to be derived.
- In principle, this involves the following aspects
  - Cholesky decomposition of latent factors:

$$\xi = \mathbf{A}\mathbf{z}. \quad (21)$$

- Decomposition of  $\mathbf{z}$  in two components:

$$\mathbf{z} = \begin{bmatrix} \mathbf{z}_1 \\ \mathbf{z}_2 \end{bmatrix} = \begin{bmatrix} \mathbf{z}_1 \\ \mathbf{0} \end{bmatrix} + \begin{bmatrix} \mathbf{0} \\ \mathbf{z}_2 \end{bmatrix} \quad (22)$$

only  $\mathbf{z}_1$  is involved in nonlinear terms.

# LMS: Background

- The conditional means and covariances are then derived (we skip the details here). The main point is that both are functions of  $z_1$

$$\mu(\mathbf{z}_1) = \begin{bmatrix} \mu_x(\mathbf{z}_1) \\ \mu_y(\mathbf{z}_1) \end{bmatrix} \quad (23)$$

and

$$\Sigma(\mathbf{z}_1) = \begin{bmatrix} \Sigma_{xx}(\mathbf{z}_1) & \Sigma_{xy}(\mathbf{z}_1) \\ \Sigma'_{xy}(\mathbf{z}_1) & \Sigma_{yy}(\mathbf{z}_1) \end{bmatrix}. \quad (24)$$

- The observed variables are then conditional normal distributed:

$$[\mathbf{x}, \mathbf{y} | \mathbf{z}_1 = z] \sim MVN(\mu(\mathbf{z}_1), \Sigma(\mathbf{z}_1)) \quad (25)$$

# LMS: Background

- These conditional means and covariances are then used for the density function

$$f(\mathbf{x} = x, \mathbf{y} = y) = \int_{\mathbf{R}^k} \varphi_{\mathbf{0}, \mathbf{I}}(\mathbf{z}_1) \varphi_{\mu(\mathbf{z}_1), \Sigma(\mathbf{z}_1)}(x, y) d\mathbf{z}_1. \quad (26)$$

where the  $\varphi_{\mathbf{0}, \mathbf{I}}(\mathbf{z}_1)$  is a multivariate standard normal density and  $\varphi_{\mu(\mathbf{z}_1), \Sigma(\mathbf{z}_1)}(x, y)$  is a conditional multivariate distribution.

- This distribution can be viewed as continuous mixture distribution with mixing weights  $\mathbf{z}_1$ .

# LMS: Background

- This mixture distribution cannot be solved analytically. However, it can be approximated with a finite mixture distribution using a Hermite-Gauss-quadrature

$$f(\mathbf{x} = x, \mathbf{y} = y) \approx \sum_{m=1}^M \pi_m \varphi_{\mu(2^{1/2}\nu_{\mathbf{m}}), \Sigma(2^{1/2}\nu_{\mathbf{m}})}(x, y). \quad (27)$$

where  $m = 1, \dots, M$  are nodes and  $\pi_m$  are mixing weights.

- Intuitive idea: the standard normal  $\mathbf{z}_1$  is replaced by a grouping variable where scores are grouped in intervals defined by the nodes.
- The estimation of finite mixtures will be part of tomorrow's workshop slides (EM algorithm).

# QML: Background

- The main difference between QML (Klein & Muthén, 2007) and LMS is the approximation of the nonnormal density function. In QML no mixture is used but instead a product of a multivariate normal density function  $f_2$  and a univariate conditionally normal density  $f_3^*$

$$\begin{aligned} f(x_i, y_i) &= f_2(x_i, \mathbf{R}y_i) f_3(y_{1i}|x_i, \mathbf{R}y_i) \\ &\approx f_2(x_i, \mathbf{R}y_i) f_3^*(y_{1i}|x_i, \mathbf{R}y_i) =: f^*(x_i, y_i), \end{aligned} \quad (28)$$

with conditional mean  $E[y_{1i}|x_i, \mathbf{R}y_i]$  and variance  $Var(y_{1i}|x_i, \mathbf{R}y_i)$ .

- Parameters are estimated by maximizing the quasi-loglikelihood function for the density  $f^*(x_i, y_i)$  using standard numerical methods (e.g., Newton-Raphson).
- The main advantages are: QML is faster, and theoretically more robust to nonnormality

## 2SMM: Background

- The 2 stage method of moments approach uses a completely different approach:
  - ① Calculate factor scores using a cfa model
  - ② Calculate the structural model using a corrected minimization criterion (very similar to OLS regression)

$$\hat{\gamma} = \hat{M}^{-1} \hat{m}, \quad (29)$$

where  $\gamma$  include the regression coefficients.

- $\hat{M}$  : corrected sum of squares and cross-products of the predictor variables (equal to  $X'X$  in a regression framework)
- $\hat{m}$  : corrected cross-products between the predictor and dependent variables (equal to  $X'y$  in a regression framework).

It is a so called error-in-variables regression that takes the unreliability of the factor score estimation from the first stage into account for the second stage.

# Empirical example 1

- The dataset was taken from PISA 2009 (OECD, 2010)
- Australian subsample of  $N = 1,019$  students who took part in a reading test.
- Constructs:
  - *Predictor 1*: Students' attitude towards reading (*Att*)
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# Summary

- Original approach used product indicators
- More recent approaches (LMS, QML, 2SMM) use different operationalization of latent product terms and provide more efficient and/or more robust estimators
- There are more approaches available, however they are typically not available in standard software or have been shown to be inefficient
- All approaches are somewhat sensitive to nonnormal distributions and assume specific functional forms of nonnormality
- Tomorrow, we will address these issues using
  - Mixture models
  - Discuss options for model fit

# Results from simulation studies (Brandt, Kelava, & Klein, 2014; Kelava, Nagengast, &

Brandt, 2014)

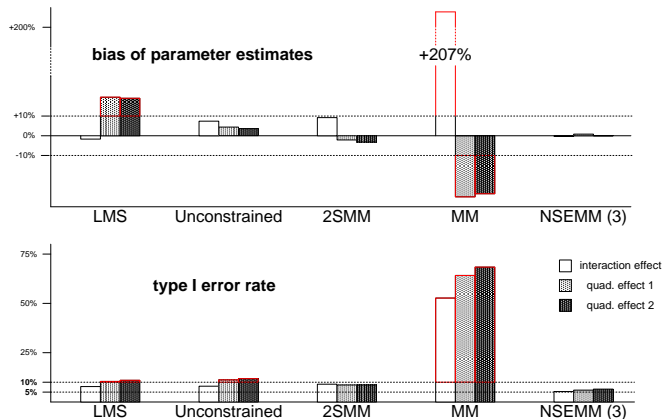


Figure: Bias of parameter estimation (top panel) and type I error rate (bottom panel) under the condition of nonnormally distributed data

Thank you for your attention.

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