## Linear Regression

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## Outline

(1) Package Check!
(2) Check the Data

- read.table plus
- Recodes
(3) One-Predictor Linear Regression
- The $\operatorname{Im}()$ function and $R$ formula
- Access Points
- About Formulas
- Diagnostics
- The Predicted Value Framework
- Categorical Predictors
- Bug-Shooting

4 Add More Predictors

- Formulas
- Moderator = categorical interaction


## Outline

- Multi-Category factor
- Numerical Interaction
(5) Conclusion


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3 One-Predictor Linear Regression

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Add More Predictors

- Formulas
- Moderator $=$ categorical interaction
- Multi-Category factor


## Check your packages

- Recall that the R (R Core Team, 2017) packages "stats" "graphics" "datasets" "base" "utils" and "grDevices" are loaded by default.

```
sessionInfo()
```

```
R version 3.6.0 (2019-04-26)
Platform: x86_64-pc-linux-gnu (64-bit)
Running under: Ubuntu 19.04
Matrix products: default
BLAS: /usr/lib/x86_64-linux-gnu/atlas/libblas.so.3.10.3
LAPACK: /usr/lib/x86_64-linux-gnu/atlas/liblapack.so.3.10.3
locale:
    [1] LC_CTYPE=en_US.UTF-8 LC_NUMERIC=C
        LC_TIME=en_US.UTF-8
    [4] LC_COLLATE=en_US.UTF-8
        LC_MESSAGES=en_US.UTF-8
    [7] LC_PAPER=en_US.UTF-8 LC_NAME=C LC_ADDRESS=C
[10] LC_TELEPHONE=C
        LC_IDENTIFICATION=C
    LC_MEASUREMENT=en_US.UTF-8
attached base packages:
```


## Check your packages ...

```
[1] stats graphics grDevices utils datasets methods base
loaded via a namespace (and not attached):
[1] compiler_3.6.0 tools_3.6.0
```


## Check your packages

- We'll use some addons today

If you don't already have these R packages, install them on your computer install.packages(c("car", "lmtest", "rockchalk"))

## Don't forget to check documentation

You can browse a list of all functions in a particular package (e.g., rockchalk)

```
library(rockchalk)
help(package = rockchalk)
```

or look up a help page for a specific function
?plotSlopes

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## Got Data?

- The example data is saved in "data/affect.dat"
- Unusually, this data does not have column names in row 1.

```
dat <- read.table("data/affect.dat", header =
    FALSE)
colnames(dat) <- c("Agency1", "Agency2",
    "Agency3",
    "Intrin1", "Intrin2", "Intrin3",
    "Extrin1", "Extrin2", "Extrin3",
    "PosAFF1", "PosAFF2", "PosAFF3",
    "NegAFF1", "NegAFF2", "NegAFF3",
    "Sex", "Ethnic2", "Ethnic3",
        "Ethnic4")
```

- View first few rows of data

$$
\begin{aligned}
& \text { options("width" }=70 \text { ) } \\
& \text { head(dat) }
\end{aligned}
$$

## Got Data?

|  |  | Agency 1 | Agency 2 | Agency 3 | Intrin1 | Intrin2 | Intrin3 | Extrin 1 | Extrin2 |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | 1 | 3.5000 | 4.0000 | 4.0000 | 4.0000 | 4.0 | 4 | 1.0000 | 1.0000 |
|  | 2 | 2.5000 | 3.1667 | 3.0000 | 3.2123 | 2.0 | 3 | 1.8333 | 2.6667 |
|  | 3 | 1.8333 | 2.0000 | 1.5000 | 3.0000 | 3.0 | 2 | 1.0000 | 1.0000 |
| 5 | 4 | 2.7714 | 3.0602 | 2.3639 | 3.1337 | 4.0 | 3 | 1.0774 | 1.1667 |
|  | 5 | 3.1667 | 3.3333 | 2.8333 | 3.5000 | 4.0 | 4 | 1.8333 | 2.0000 |
|  | 6 | 2.3333 | 2.8333 | 2.3333 | 3.0000 | 2.5 | 3 | 3.0588 | 2.4125 |
|  |  | $\begin{array}{r} \text { Extrin3 } \\ \text { Eth } \end{array}$ | $\begin{aligned} & \text { PosAFF1 } \\ & \text { nic2 } \end{aligned}$ | PosAFF2 | PosAFF3 | NegAFF1 | NegAFF2 | NegAFF3 | Sex |
|  | 1 | 1.5000 | $\begin{aligned} & 4.0000 \\ & 0 \end{aligned}$ | 4.0 | 4.0 | 1.0 | 1.0000 | 1.0 | 1 |
| 10 | 2 | 1.8333 | $\begin{aligned} & 3.0000 \\ & 0 \end{aligned}$ | 3.5 | 2.5 | 1.5 | 1.6858 | 1.5 | 1 |
|  | 3 | 1.0000 | $\begin{aligned} & 3.0184 \\ & 0 \end{aligned}$ | 2.5 | 3.0 | 1.0 | 1.0000 | 1.0 | 1 |
|  | 4 | 1.0000 | $\begin{aligned} & 3.0000 \\ & 0 \end{aligned}$ | 2.5 | 3.0 | 2.5 | 2.5000 | 1.5 | 1 |
|  | 5 | 1.8333 | $\begin{aligned} & 3.7804 \\ & 0 \end{aligned}$ | 3.5 | 3.0 | 2.5 | 2.0000 | 3.0 | 1 |
|  | 6 | 2.6667 | $\begin{aligned} & 4.0000 \\ & 0 \end{aligned}$ | 3.0 | 3.0 | 2.0 | 1.5000 | 2.0 | 1 |
| 15 |  | Ethnic3 | Ethnic4 |  |  |  |  |  |  |
|  | 1 | 1 | 0 |  |  |  |  |  |  |
|  | 2 | 0 | 0 |  |  |  |  |  |  |
|  | 3 | 0 | 0 |  |  |  |  |  |  |

## Got Data?

$20 |$| 4 | 0 | 0 |
| :--- | :--- | :--- |
| 5 | 0 | 0 |
| 6 | 0 | 0 |

options("width" = 80)

## recodes

- Create scales by calculating means of the indicator variables

```
dat$agency <- rowMeans(dat [ ,
    c("Agency1","Agency2","Agency3")], na.rm =
    TRUE)
dat$intMotiv <- rowMeans(dat[ ,
    c("Intrin1","Intrin2","Intrin3")], na.rm =
    TRUE)
dat$extMotiv <- rowMeans(dat [ ,
    c("Extrin1","Extrin2","Extrin3")], na.rm =
    TRUE)
dat$posAffect <- rowMeans(dat[ ,
    c("PosAFF1","PosAFF2","PosAFF3")], na.rm =
    TRUE)
dat$negAffect <- rowMeans(dat[ ,
    c("NegAFF1","NegAFF2","NegAFF3")], na.rm =
    TRUE)
```


## recodes ...

- Recode dummy variables
table(dat\$Sex)

| 1 | 2 |
| ---: | ---: |
| 195 | 185 |

```
dat\$gender <- factor (dat\$Sex, levels \(=c(1,2)\),
    labels = c("male", "female"))
dat\$ethnicity <- as.factor (ifelse (dat\$Ethnic2,
    "Black",
    ifelse (dat\$Ethnic3,
    "Hispanic",
ifelse (dat\$Ethnic4,
    "Asian",
    "White")))
```


## recodes ...

```
dat$race <-
    rockchalk::combineLevels(dat$ethnicity, levs
    = c("Black", "Hispanic", "Asian"), newLabel =
    "Nonwhite")
```

```
The original levels Asian Black Hispanic White
have been replaced by White Nonwhite
```

options("width" = 60)
head (dat)

## recodes ...

|  | Agency 1 | Agency 2 | Agency 3 | Intrin 1 | Intrin2 | Intrin3 | Extrin 1 |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 1 | 3.5000 | 4.0000 | 4.0000 | 4.0000 | 4.0 | 4 | 1.0000 |
| 2 | 2.5000 | 3.1667 | 3.0000 | 3.2123 | 2.0 | 3 | 1.8333 |
| 3 | 1.8333 | 2.0000 | 1.5000 | 3.0000 | 3.0 | 2 | 1.0000 |
| 4 | 2.7714 | 3.0602 | 2.3639 | 3.1337 | 4.0 | 3 | 1.0774 |
| 5 | 3.1667 | 3.3333 | 2.8333 | 3.5000 | 4.0 | 4 | 1.8333 |
| 6 | 2.3333 | 2.8333 | 2.3333 | 3.0000 | 2.5 | 3 | 3.0588 |
|  | Extrin2 | Extrin3 | PosAFF1 | PosAFF2 | PosAFF3 | NegAFF1 | NegAFF2 |
| 1 | 1.0000 | 1.5000 | 4.0000 | 4.0 | 4.0 | 1.0 | 1.0000 |
| 2 | 2.6667 | 1.8333 | 3.0000 | 3.5 | 2.5 | 1.5 | 1.6858 |
| 3 | 1. 0000 | 1.0000 | 3. 0184 | 2.5 | 3.0 | 1.0 | 1.0000 |
| 4 | 1. 1667 | 1.0000 | 3.0000 | 2.5 | 3.0 | 2.5 | 2.5000 |
| 5 | 2.0000 | 1.8333 | 3.7804 | 3.5 | 3.0 | 2.5 | 2.0000 |
| 6 | 2.4125 | 2.6667 | 4.0000 | 3.0 | 3.0 | 2.0 | 1.5000 |
|  | NegAFF3 | Sex Eth | c2 Ethn | ic3 Eth | c 4 | gency in | Motiv |
| 1 | 1.0 | 1 | 0 | 1 | 03. | 333334 | 00000 |
| 2 | 1.5 | 1 | 0 | 0 | 02.8 | 889002. | 37433 |
| 3 | 1.0 | 1 | 0 | 0 | 01. | 77767 2. | 66667 |
| 4 | 1.5 | 1 | 0 | 0 | 02.7 | 31833 3. | 77900 |
| 5 | 3.0 | 1 | 0 | 0 | 03.1 | 111003. | 833333 |
| 6 | 2.0 | 1 | 0 | 0 | 02.4 | 99672.8 | 833333 |
|  | extMotiv | posAffe | ct negA | $f e c t$ ge | nder eth | icity | race |
| 1 | 1.166667 | 4.0000 | 0001.00 | 00000 | male Hi | panic | nwhite |
| 2 | 2.111100 | 3.0000 | 001.56 | 61933 | male | White | White |
| 3 | 1.000000 | 2.8394 | $67 \quad 1.0$ | 00000 | male | White | White |

## recodes ...

$|$| 4 | 1.081367 | 2.833333 | 2.166667 | male | White | White |
| :--- | :--- | :--- | :--- | :--- | :--- | :--- |
| 5 | 1.888867 | 3.426800 | 2.500000 | male | White | White |
| 6 | 2.712667 | 3.333333 | 1.833333 | male | White | White |

options("width" = 80)

- Save a copy of that in the workingdata folder saveRDS(dat, file = "workingdata/affect.rds")


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## R formula

- Almost all model fitting functions in R use the Wilkinson and Rogers notation

$$
\text { dependent_variable } \sim \text { predictor_variable }
$$

- Can omit the estimation of the intercept
- Old fashioned way

$$
\text { dependent_variable } \sim-1+\text { predictor_variable }
$$

- The old fashioned way confused school children, hence the new fashioned way

$$
\text { dependent_variable } \sim 0 \text { + predictor_variable }
$$

- There are special symbols in R formula notation, "+",","," "*", "/", "^", "|".


## The $\operatorname{Im}()$ function

- Suppose we want to explain Positive Affect with sense of Agency

```
reg.mod.1 <- lm(posAffect ~ agency, data = dat)
```

Returns "silently" unless there is an error

- Print method for Im objects offers minimal information
reg.mod. 1

```
Call:
lm(formula = posAffect ~ agency, data = dat)
Coefficients:
(Intercept) agency
    2.1883 0.3491
```

- There is quite a bit of structure in there, however. Run "str()" you will see. I'll run the briefer "attributes".


## The $\operatorname{Im}()$ function

attributes(reg.mod.1)

```
$names
\begin{tabular}{llll}
{\([1]\)} & "coefficients" "residuals" & "effects" & "rank" \\
{\([5]\)} & fitted.values" "assign" & "qr" & "df.residual" \\
{\([9]\)} & "xlevels" & "call" & "terms"
\end{tabular}
$class
[1] " lm"
```


## Direct Retrieval versus Accessor functions

- The Im object "reg.mod.1" is of class "Im", S3 object.
class(reg.mod.1)
[1] "lm"
- If you ran "str(reg.mod.1)", a big structure inside there would be apparent.
- I'll just ask for names

```
names(reg.mod.1)
```

```
[1] "coefficients" "residuals" "effects"
[4] "rank" "fitted.values" "assign"
[7] "qr" "df.residual" "xlevels"
[10] "call" "terms" "model"
```


## Everybody Loves \$

- S3 list objects allow a shortcut access with the dollar sign
- data.frame access like dat $\$ \times 1$
- Notice that inside the fitted model object there is an element named "df.residual". Get that:

```
reg.mod.1$df.residual
```

```
[1] 378
```

Since the dollar sign is a shortcut notation, we could go the long form as well

```
reg.mod.1[["df.residual"]]
```


## [1] 378

- Any of the elements in reg.mod.1's internal structure can be retrieved in that way.


## Everybody Loves \$ ...

- If this were an S4 class object, then we would use the "@" sign rather than the " $\$$ " sign as a shortcut.
- The R Core team does NOT encourage us to pull pieces out in that way.
- They reserve the right to rename those internal bits.
- Instead, it is recommended to use "accessor" functions that R provides


## Coefficients, retrieved both ways

- Point estimates of parameters (regression coefficients)
(1) The accessor function "coef()" (short for coefficients)

$$
\operatorname{coef}(\mathrm{reg} . \bmod .1)
$$

```
(Intercept) agency
    2.1882796 0.3491155
```

2 Use the dollar sign access

```
reg.mod.1$coefficients
```

```
(Intercept)
    agency
    2.1882796 0.3491155
```

- The predicted value vector
(1) The accessor function "fitted()"
head (fitted (reg.mod.1))


## Coefficients, retrieved both ways ...

| 1 | 2 | 3 | 4 | 5 | 6 |
| ---: | ---: | ---: | ---: | ---: | ---: |
| 3.526556 | 3.196839 | 2.808925 | 3.142005 | 3.274413 | 3.061057 |

(2) The dollar sign avenue (note element name "fitted.values" is different than accessor function name "fitted()")

```
head(reg.mod.1$fitted.values)
```

| 1 | 2 | 3 | 4 | 5 | 6 |
| ---: | ---: | ---: | ---: | ---: | ---: |
| 3.526556 | 3.196839 | 2.808925 | 3.142005 | 3.274413 | 3.061057 |

## Coefficients, retrieved both ways ...

- The residual vector
(1) The accessor function "resid()"
head(resid(reg.mod.1), 4)

| 1 | 2 | 3 | 4 |
| ---: | ---: | ---: | ---: |
| 0.47344443 | -0.19683927 | 0.03054124 | -0.30867153 |

(2) The dollar sign avenue (note element name is different than accessor function)

```
head(reg.mod.1$residuals, 4)
```

    1 2 3 4
    $0.47344443-0.19683927 \quad 0.03054124-0.30867153$

## Some functions offer much more elaborate information

- Every useful object in R is supposed to have a summary () method!
- The " summary () " function is as close as we get to a "big standard output"
summary (reg.mod.1)

```
Call:
lm(formula = posAffect ~ agency, data = dat)
Residuals:
\begin{tabular}{rrrrr} 
Min & \(1 Q\) & Median & 3Q & Max \\
-2.06107 & -0.37515 & 0.04591 & 0.45144 & 1.39929
\end{tabular}
Coefficients:
            Estimate Std. Error t value Pr (>|t|)
l(Intercept) 2.18828 
Signif. codes:
0,****' 0.001 ,**' 0.01 '*'0.05 ,., 0.1 , , 1
Residual standard error: 0.6077 on 378 degrees of freedom
```


## Some functions offer much more elaborate information ...

```
Multiple R-squared: 0.07664, Adjusted R-squared: 0.0742
F-statistic: 31.37 on 1 and 378 DF, p-value: 4.103e-08
```

- Confidence intervals for regression coefficients

```
confint(reg.mod.1, level = .99)
```

```
    0.5 % 99.5 %
(Intercept) 1.7750122 2.6015469
agency 0.1877538}0.510477
```


## Math functions in Formulas

- If we wanted to predict with the logarithm of a variable,
(1) We could create a new variable by recoding, Or
(2) use the symbol for the logarithm in the formula

$$
\text { dependent_variable } \sim \log (p r e d i c t o r)
$$

Any of the mathematical transformations in R could be used in place of log.
dependent_variable ~ sqrt(predictor)

- I don't usually write math transformations into formulas, it complicates plotting and table-making duties later on.


## Special Symbols in Formulas

- Multiple regression: "+" for more predictors

```
dependent_variable ~ predictor1 + predictor2
```

- Interaction: "*"
dependent_variable ~ predictor1 * predictor2
Means you want a regression to estimate 3 coefficients, $\beta_{1}$ predictor $1+\beta_{2}$ predictor $2+\beta_{3}$ predictor $1 \times$ predictor 2


## Be cautious with

- You may think to yourself, "I'll add a squared term":

```
dependent_variable ~ predictor + predictor^2
```

- However, there is a gotcha
(1) "^" has a special meaning in the formula notation.
(2) If we are trying to make a predictive equation like

$$
\text { dependent_variable }=\beta_{0}+\beta_{1} \text { predictor }+\beta_{2} \text { predictor }^{2}
$$

Wrap the math inside the capital $\mathbf{I}()$.

$$
\begin{gathered}
\text { dependent_variable } \sim \text { predictor } \\
+\quad I\left(\text { predictor }{ }^{\wedge} 2\right)
\end{gathered}
$$

(3) But don't do that, better ways exist (orthogonal polynomials).

## plot diagnostics

The Im class has a plot method (plot.Im )
plot(reg.mod.1)
defaults to offer 4 graphs (can be adjusted, see ? plot.Im )

## plot diagnostics ...



## plot diagnostics ...

Plot diagnostics

1. residuals are (not) related to fitted values
2. residuals are (not) approximately normally distributed
3. residuals are (not) homoscedastic
4. highly influential (leverage) observations do/not exist (Cook's Distance)

## Influence Diagnostics

- The influence.measures() function collects a great deal of information and displays information for each row in the data for the fitted model:

```
inf1 <- influence.measures(reg.mod.1)
```

- Output is immense, does not fit within these notes


## inf 1

generates such a massive outflow that the presentation software fails.

- A tidy summary function show the problematic cases
summary (inf1)


## Influence Diagnostics

|  | ntial | influe | al ob | vatio |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | 1 m (for | $\mathrm{a}=\mathrm{po}$ | fect | ency | data | dat |
|  | dfb. 1 | dfb.agnc | dffit | cov.r | cook.d | hat |
| 1 | -0.10 | 0.11 | 0.12 | 1.02_* | 0.01 | 0.02 * |
| 25 | -0.02 | 0.03 | 0.03 | 1.02_* | 0.00 | 0.02 |
| 26 | 0.31 | -0.30 | 0.31 -* | 1.01 | 0.05 | 0.03_* |
| 38 | 0.07 | -0.06 | 0.07 | 1.02_* | 0.00 | 0.01 |
| 95 | 0.11 | -0.12 | -0.13 | 1.02_* | 0.01 | 0.02_* |
| 129 | -0.06 | 0.07 | 0.08 | 1.02_* | 0.00 | 0.01 |
| 131 | 0.08 | -0.09 | -0.10 | 1.02_* | 0.01 | 0.02 * |
| 136 | 0.14 | -0.13 | 0.14 | 1.03_* | 0.01 | 0.03_* |
| 146 | -0.10 | 0.11 | 0.11 | 1.03_* | 0.01 | 0.03_* |
| 177 | -0.02 | -0.01 | -0.16 | 0.96 * | 0.01 | 0.00 |
| 196 | 0.02 | -0.04 | -0.13 | 0.98 * | 0.01 | 0.00 |
| 211 | -0.06 | 0.04 | -0.12 | 0.98_* | 0.01 | 0.00 |
| 229 | 0.27 | -0.26 | 0.28_* | 1.00 | 0.04 | 0.02 * |
| 230 | 0.07 | -0.08 | -0.08 | 1.02_* | 0.00 | 0.02_* |
| 232 | -0.16 | 0.14 | -0.20 | 0.97 _* | 0.02 | 0.00 |
| 245 | 0.11 | -0.10 | 0.11 | 1.02 * | 0.01 | 0.02 * |
| 252 | -0.06 | 0.07 | 0.07 | 1.03_* | 0.00 | 0.02_* |
| 256 | 0.04 | -0.05 | -0.06 | 1.02_* | 0.00 | 0.01 |
| 261 | -0.03 | 0.03 | 0.04 | 1.02_* | 0.00 | 0.01 |
| 280 | 0.05 | -0.05 | 0.05 | 1.03_* | 0.00 | 0.02_* |
| 294 | 0.05 | -0.05 | 0.05 | 1.02_* | 0.00 | 0.01 |

## Influence Diagnostics

| 305 | 0.34 | -0.32 | $0.35_{-} *^{*}$ | 1.00 | 0.06 | $0.02_{-}^{*}$ |
| ---: | ---: | ---: | :--- | :--- | :--- | :--- |
| 336 | -0.04 | 0.00 | -0.18 | $0.95_{-}{ }^{*}$ | 0.02 | 0.00 |
| 353 | 0.05 | -0.05 | -0.06 | $1.02_{-}{ }^{*}$ | 0.00 | 0.02 |
| 368 | 0.18 | -0.21 | $-0.25_{-}{ }^{*}$ | $0.98_{-}{ }^{*}$ | 0.03 | 0.01 |
| 373 | 0.01 | -0.01 | -0.01 | $1.02_{-}^{*}$ | 0.00 | 0.01 |

## Thumbnail sketch

- In case you wonder what those things are, I wrote out a longer lecture about it http://pj.freefaculty.org/guides/stat/Regression/ RegressionDiagnostics

| dfb.1_ | dfb.agnc | dffit | cov.r | cook.d | hat |
| ---: | ---: | ---: | ---: | ---: | ---: |
| -0.098 | 0.108 | 0.115 | 1.024 | 0.007 | 0.021 |

- Thumbnail sketch is as follows
dfb "df beta" (one for each coefficient) shows how the coefficient estimate would change if that row were dropped
dffit change in predicted value if that row were dropped cook.d A summary of how far the vector of parameter estimates ( $\hat{\beta}_{0}, \hat{\beta}_{1}$ ) would change if that row were dropped.
hat The "Hat matrix" value for the row. If this value is large, it means the case is influential in changing the overall regression line


## The predict function accepts a "newdata" argument

- predict(reg.mod.1) is the same result as fitted(reg.mod.1) : the predicted values of the observed cases.
- Often, we want predicted values for particular, substantively interesting values of the predictors.
- Obtain predicted values for a new data set:

```
predict(reg.mod.1, newdata =
    some_data_frame_we_make_up)
```

- And if we want confidence intervals, we add

```
predict(reg.mod.1, newdata =
    some_data_frame_we_make_up, interval =
    "confidence")
```

- This makes it possible to calculate "marginal effects", the change in the outcome due to any given change in a predictor.


## rockchalk::newdata

- The newdata object MUST include
- all predictors, with exactly same names as used in the formula, and
- values of factors within the newdata object must match the data used to fit the model
- In rockchalk, I needed this often and wrote a "newdata" function.
- For example, I notice the variable "agency" varies between 1 and 4 .

```
library(rockchalk)
nd <- newdata(reg.mod.1, predVals =
    c("agency"), n = 5)
nd
```

```
    agency
11.000000
2 2.175525
3 2.499983
4 2.832975
54.000000
```


## rockchalk::newdata ...

```
## n = 5 is 5 evenly spaced quartile values
```

- Then we use that with the predict function

```
reg.mod.1.pred \(<-\) predict (reg.mod.1, newdata
    = nd)
reg.mod.1.pred
```

| 1 | 2 | 3 | 4 | 5 |
| ---: | ---: | ---: | ---: | ---: |
| 2.537395 | 2.947789 | 3.061062 | 3.177315 | 3.584741 |

## Stash Predictions into the newdata frame

- Usually, you want to save the fitted values

```
nd$reg.mod.1.pred <- predict(reg.mod.1, newdata =
    nd)
```

- Because I became tired of that, in rockchalk I wrote predictOMatic(). It creates the new data and also saves the predictions:

```
reg.mod.1.pm <- predict0Matic(reg.mod.1, predVals
    = c("agency"), n = 5)
reg.mod.1.pm
```

```
    agency fit
1 1.000000 2.537395
2 2.175525 2.947789
3 2.499983 3.061062
4 2.832975 3.177315
54.000000 3.584741
```


## Stash Predictions into the newdata frame ...

- A more-or-less "automatic" graphing routine, " plotSlopes ", will do all of this and draw a plot. Before I show that, I need to show about confidence intervals for predictions


## Confidence Interval on Predicted Values

- The R predict() function has a confidence interval argument. It defaults to none, but can be either "confidence" or "prediction".
- The returned data structure is a matrix with 3 columns

```
reg.mod.1.pred2 <- predict(reg.mod.1, newdata =
    nd, interval = "confidence")
reg.mod.1.pred2
```

|  | fit | lwr | upr |
| ---: | ---: | ---: | ---: |
| 1 | 2.537395 | 2.342241 | 2.732549 |
| 2 | 2.947789 | 2.873926 | 3.021652 |
| 3 | 3.061062 | 2.999750 | 3.122375 |
| 4 | 3.177315 | 3.104471 | 3.250159 |
| 5 | 3.584741 | 3.392334 | 3.777149 |

- For the 5 example values of agency, we have a value
(1) "on" the line, and
(2) $95 \%$ range below ("lwr")

3 $95 \%$ range above ("upr")

## The CI lines should be a "smooth hourglass"



Those CI lines are "connect-the-dots" curves. In this case, they don't look so bad

## plotSlopes with no confidence interval

plotSlopes(reg.mod.1, plotx = "agency")


## predictOMatic understands the interval argument

```
reg.mod.1.pm <- predictOMatic (reg.mod.1, predVals
    \(=c(" a g e n c y "), n=5\), interval = "confidence")
reg.mod.1.pm
```

|  | agency | fit | lwr | upr |
| ---: | ---: | ---: | ---: | ---: |
| 1 | 1.000000 | 2.537395 | 2.342241 | 2.732549 |
| 2 | 2.175525 | 2.947789 | 2.873926 | 3.021652 |
| 3 | 2.499983 | 3.061062 | 2.999750 | 3.122375 |
| 4 | 2.832975 | 3.177315 | 3.104471 | 3.250159 |
| 5 | 4.000000 | 3.584741 | 3.392334 | 3.777149 |

plotSlopes (reg.mod.1, plotx $=$ "agency", $n=5$, interval = "confidence")

## predictOMatic understands the interval argument ...



- plotSlopes creates an output object that has the newdata in it:
reg.mod.ps <- plotSlopes(reg.mod.1, plotx =

$$
\text { "agency", } n=5 \text {, interval = "confidence") }
$$

## predictOMatic understands the interval argument ...

- To obtain smooth curve, I need to calculate predictions for more values of $X$. plotSlopes calculates predicted values for, $\mathrm{n}=40$, values of agency
reg.mod.ps\$newdata

|  | agency | fit | lwr | upr |
| :--- | ---: | ---: | ---: | ---: |
| 1 | 1.000000 | 2.537395 | 2.342241 | 2.732549 |
| 2 | 1.076923 | 2.564250 | 2.378022 | 2.750478 |
| 3 | 1.153846 | 2.591105 | 2.413752 | 2.768458 |
| 4 | 1.230769 | 2.617960 | 2.449422 | 2.786498 |
| 5 | 1.307692 | 2.644815 | 2.485022 | 2.804608 |
| 6 | 1.384615 | 2.671670 | 2.520540 | 2.822801 |
| 7 | 1.461538 | 2.698525 | 2.555961 | 2.841090 |
| 8 | 1.538462 | 2.725380 | 2.591266 | 2.859495 |
| 9 | 1.615385 | 2.752235 | 2.626432 | 2.878039 |
| 10 | 1.692308 | 2.779090 | 2.661430 | 2.896751 |
| 11 | 1.769231 | 2.805945 | 2.696222 | 2.915669 |
| 12 | 1.846154 | 2.832800 | 2.730760 | 2.934841 |
| 13 | 1.923077 | 2.859655 | 2.764982 | 2.954329 |
| 14 | 2.000000 | 2.886511 | 2.798809 | 2.974212 |
| 15 | 2.076923 | 2.913366 | 2.832139 | 2.994592 |
| 16 | 2.153846 | 2.940221 | 2.864843 | 3.015598 |

## predictOMatic understands the interval argument ...

| 17 | 2.230769 | 2.967076 | 2.896766 | 3.037385 |
| :--- | :--- | :--- | :--- | :--- |
| 18 | 2.307692 | 2.993931 | 2.927727 | 3.060134 |
| 19 | 2.384615 | 3.020786 | 2.957539 | 3.084032 |
| 20 | 2.461538 | 3.047641 | 2.986036 | 3.109245 |
| 21 | 2.538462 | 3.074496 | 3.013113 | 3.135878 |
| 22 | 2.615385 | 3.101351 | 3.038755 | 3.163947 |
| 23 | 2.692308 | 3.128206 | 3.063041 | 3.193371 |
| 24 | 2.769231 | 3.155061 | 3.086123 | 3.223998 |
| 25 | 2.846154 | 3.181916 | 3.108186 | 3.255646 |
| 26 | 2.923077 | 3.208771 | 3.129414 | 3.288128 |
| 27 | 3.000000 | 3.235626 | 3.149972 | 3.321280 |
| 28 | 3.076923 | 3.262481 | 3.169996 | 3.354966 |
| 29 | 3.153846 | 3.289336 | 3.189595 | 3.389077 |
| 30 | 3.230769 | 3.316191 | 3.208857 | 3.423525 |
| 31 | 3.307692 | 3.343046 | 3.227847 | 3.458245 |
| 32 | 3.384615 | 3.369901 | 3.246618 | 3.493184 |
| 33 | 3.461538 | 3.396756 | 3.265210 | 3.528302 |
| 34 | 3.538462 | 3.423611 | 3.283655 | 3.563567 |
| 35 | 3.615385 | 3.450466 | 3.301978 | 3.598955 |
| 36 | 3.692308 | 3.477321 | 3.320198 | 3.634444 |
| 37 | 3.769231 | 3.504176 | 3.338332 | 3.670020 |
| 38 | 3.846154 | 3.531031 | 3.356393 | 3.705670 |
| 39 | 3.923077 | 3.557886 | 3.374391 | 3.741382 |
| 40 | 4.000000 | 3.584741 | 3.392334 | 3.777149 |

## R factors are recoded into "dummy variables" in regression models

- gender is a dichotomous variable, coded "male" or "female". Check the levels:
levels(dat\$gender)

```
[1] "male" "female"
```


## R factors are recoded into "dummy variables" in regression models ...

## - Include gender as a predictor

```
reg.mod.2 <- lm(posAffect ~ gender, data = dat)
summary(reg.mod.2)
```

```
Call:
lm(formula = posAffect ~ gender, data = dat)
Residuals:
    Min 1Q Median 3Q Max
-2.10992 -0.35609 -0.01197 0.47724 0.97724
```

Coefficients:

|  | Estimate | Std. Error | t value | Pr $(>\|t\|)$ |  |
| :--- | ---: | ---: | ---: | ---: | ---: |
| (Intercept) | 3.02276 | 0.04518 | 66.906 | $<2 e-16$ | $* * *$ |
| genderfemale | 0.08717 | 0.06475 | 1.346 | 0.179 |  |

Signif. codes:
$0^{\prime 2 * *} 0.001^{\prime * *} 0.01$,*, 0.05 ,.' 0.1 , 1
Residual standard error: 0.6309 on 378 degrees of freedom
Multiple R-squared: 0.004772, Adjusted R-squared: 0.002139
F-statistic: 1.812 on 1 and 378 DF, p-value: 0.179

## R factors are recoded into "dummy variables" in regression models ...

- One gender, in this case "males", is treated as a baseline group. There are 2 categories, we can only estimate 2 coefficients. The default model include an intercept, then only 1 coefficient is left over for one of the groups.
- In this case, the predicted values would be
- males: 3.0227
- females: $3.0227+0.08717$


## R factors are recoded into "dummy variables" in regression models ...

- There are 2 alternatives to this coding scheme
(1) Get rid of the intercept, in which case we get one estimate for males, one for females

```
reg.mod.2b <- lm(posAffect ~ 0 + gender, data
    = dat)
summary(reg.mod.2b)
```

```
Call:
lm(formula = posAffect ~ 0 + gender, data = dat)
Residuals:
\begin{tabular}{rrrrr} 
Min & \(1 Q\) & Median & 3Q & Max \\
-2.10992 & -0.35609 & -0.01197 & 0.47724 & 0.97724
\end{tabular}
Coefficients:
    Estimate Std. Error t value Pr(>|t|)
gendermale }\begin{array}{lllll}{3.02276 0.04518 66.91 <2e-16 ***}
genderfemale 3.10992 0.04638 67.05 <2e-16 ***
```


# R factors are recoded into "dummy variables" in regression models ... 

```
Signif. codes: 0 '***' 0.001 '**' 0.01 '*', 0.05 '., 0.1 , , 1
Residual standard error: 0.6309 on 378 degrees of freedom
Multiple R-squared: 0.9596, Adjusted R-squared: 0.9594
F-statistic: 4486 on 2 and 378 DF, p-value: < 2.2e-16
```

The disadvantage of this coding is that we cannot directly say whether the 2 values are statistically significantly different from one another.
(2) Reverse the levels on the gender variable

```
dat$gender2 <- factor(dat$gender, levels =
    c("female", "male"))
reg.mod.2c <- lm(posAffect ~ gender2, data = dat)
summary(reg.mod.2c)
```


# R factors are recoded into "dummy variables" in regression models 

```
Call:
lm(formula = posAffect ~ gender2, data = dat)
Residuals:
```



```
-2.10992 -0.35609 -0.01197 0.47724 0.97724
Coefficients:
\begin{tabular}{lrrrrr} 
& Estimate & Std. Error & t value & \(\operatorname{Pr}(>|t|)\) & \\
(Intercept) & 3.10992 & 0.04638 & 67.047 & \(<2 \mathrm{e}-16\) & \(* * *\) \\
gender2male & -0.08717 & 0.06475 & -1.346 & 0.179 &
\end{tabular}
Signif. codes:
0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ', 1
Residual standard error: 0.6309 on 378 degrees of freedom
Multiple R-squared: 0.004772, Adjusted R-squared: 0.002139
F-statistic: 1.812 on 1 and 378 DF, p-value: 0.179
```


## Test Homogeneity of Variances

- Use the Levene test, which is in John Fox's car package.

```
library(car)
leveneTest(reg.mod.2)
```

```
Levene's Test for Homogeneity of Variance (center = median)
    Df F value Pr (>F)
group 1 2.3963 0.1225
    378
```

- This suggests we were not wrong to assume the error variances for males and females are the same.


## A multi-category factor

levels(dat\$ethnicity)
[1] "Asian" "Black" "Hispanic" "White"
reg.mod. $3<-\operatorname{lm}(p o s A f f e c t \sim$ ethnicity, data $=$ dat) summary (reg.mod.3)

```
Call:
```

lm(formula $=$ posAffect $\sim$ ethnicity, data $=$ dat)
Residuals:
Min 1Q Median 3Q Max
$-2.00610-0.40094-0.01907 \quad 0.49390 \quad 1.06360$

Coefficients:

|  | Estimate | Std. Error | t value | Pr (>\|t|) |  |
| :--- | ---: | ---: | ---: | ---: | ---: |
| (Intercept) | 2.9364 | 0.1026 | 28.628 | $<2 \mathrm{e}-16$ | $* * *$ |
| ethnicityBlack | 0.0697 | 0.1777 | 0.392 | 0.695 |  |
| ethnicityHispanic | 0.1237 | 0.1284 | 0.964 | 0.336 |  |
| ethnicityWhite | 0.1536 | 0.1099 | 1.398 | 0.163 |  |

Signif. codes:

## A multi-category factor ...

```
0 ',***' 0.001 '**', 0.01 '*'0.05 '.'0.1 ' , 1
Residual standard error: 0.6323 on 376 degrees of freedom
Multiple R-squared: 0.005664, Adjusted R-squared: -0.00227
F-statistic: 0.7139 on 3 and 376 DF, p-value: 0.5442
```


## Finding out what's going wrong

- Sometimes you'll have confusing output that you can't understand
- I often snoop on data as R sees it "inside" $\operatorname{Im}()$ :
(1) model.frame: a data.frame with output and predictors that R creates when you run Im.

```
rm1.mf <- model.frame(reg.mod.1)
head(rm1.mf)
```

5 |  | posAffect | agency |
| ---: | ---: | ---: |
| 1 | 4.000000 | 3.833333 |
| 2 | 3.000000 | 2.888900 |
| 3 | 2.839467 | 1.777767 |
| 4 | 2.833333 | 2.731833 |
| 5 | 3.426800 | 3.111100 |
| 6 | 3.333333 | 2.499967 |

(2) Suppose the regression fails, so there is no object from which to obtain a frame. No problem! Give the formula to model.frame.

## Finding out what's going wrong ...

```
rm1.mf <- model.frame(posAffect ~
    log(agency), data = dat)
head(rm1.mf)
```

|  | posAffect | log(agency) |
| ---: | ---: | ---: |
| 1 | 4.000000 | 1.3437347 |
| 2 | 3.000000 | 1.0608758 |
| 3 | 2.839467 | 0.5753579 |
| 4 | 2.833333 | 1.0049729 |
| 5 | 3.426800 | 1.1349764 |
| 6 | 3.333333 | 0.9162774 |

(3) model.matrix shows the "design matrix", the numeric columns used in estimation

```
rm1.dm <- model.matrix(reg.mod.1)
head(rm1.dm)
```


## Finding out what's going wrong ...

```
(Intercept) agency
1 3.833333
    1 2.888900
    1 1.777767
    1 2.731833
    1 3.111100
    1 2.499967
```

- This is especially revealing if there is a factor as a predictor

```
rm2.dm <- model.matrix(posAffect ~ ethnicity,
        data = dat)
head(rm2.dm)
```


## Finding out what's going wrong ...

|  | (Intercept) ethnicityBlack |  |  | ethnicityHispanic |
| :---: | :---: | :---: | :---: | :---: |
|  | 1 | 1 | 0 | 1 |
|  | 2 | 1 | 0 | 0 |
|  | 3 | 1 | 0 | 0 |
| 5 | 4 | 1 | 0 | 0 |
|  | 5 | 1 | 0 | 0 |
|  | 6 | 1 | 0 | 0 |
| ethnicityWhite |  |  |  |  |
|  | 1 |  | 0 |  |
| 10 | 2 |  | 1 |  |
|  | 3 |  | 1 |  |
|  | 4 |  | 1 |  |
|  | 5 |  | 1 |  |
|  | 6 |  | 1 |  |

## Outline

Package Check!
2 Check the Data

- read.table plus
- Recodes

One-Predictor Linear Regression

- The Im() function and R formula
- Access Points
- About Formulas
- Diagnostics
- The Predicted Value Framework
- Categorical Predictors
- Bug-Shooting

4 Add More Predictors

- Formulas
- Moderator = categorical interaction
- Multi-Category factor
- Numerical Interaction


## Addition sign "+"

- Can insert math transformations "on the fly"

$$
\begin{aligned}
& \text { dep_var } \sim \log (\text { predictor1) + sqrt(predictor2) }+ \\
& \text { predictor3 + predictor4 }
\end{aligned}
$$

but this makes creating a newdata object somewhat more complicated

- However, rockchalk::newdata() and predictOMatic can work together to avoid problems for us!


## Multiplication sign "*" is not exactly like multiplication

- A multiplicative interaction between two continuous predictors can be entered like so

$$
\text { dep_var } \sim \text { predictor } 1 * \text { predictor } 2+\text { predictor } 3
$$

- It adds predictor1 and predictor2 as "additive" (or "main") effects, plus their product.

```
dep_var ~ predictor1 + predictor2 +
    predictor1:predicor2 + predictor3
```

- COLON! The symbol "predictor1:predictor2" represents "predictor $1 \times$ predictor 2 ".


## With categorical predictors, "*" does something else

- Because factor variables are broken into dummy variables, an interactive term like

```
posAffect ~ gender * agency
```

will have the effect of estimating a different slope and a different intercept for each of the sexes. Will illustrate in next section.

## Our first guess might be "everything is additive"

```
reg.mod. 5 <- lm(posAffect ~ agency + gender, data
    = dat)
```

summary (reg.mod.5)
Call:
lm(formula $=$ posAffect $\sim$ agency + gender, data $=$ dat)
Residuals:
Min 1Q Median 3Q Max

| -2.10427 | -0.39890 | 0.05395 | 0.44156 | 1.35513 |
| :--- | :--- | :--- | :--- | :--- |

Coefficients:

| Estimate | Std. Error | t value | $\operatorname{Pr}(>\|\mathrm{t}\|)$ |  |
| ---: | ---: | ---: | ---: | ---: | ---: |
| 2.14912 | 0.16207 | 13.260 | $<2 \mathrm{e}-16$ | $* * *$ |
| 0.34839 | 0.06226 | 5.596 | $4.24 \mathrm{e}-08$ | $* * *$ |
| 0.08417 | 0.06230 | 1.351 | 0.177 |  |

Signif. codes:
0 , ***' 0.001 , **, 0.01 , ${ }^{\prime} 0.05$, , 0.1 , 1

Residual standard error: 0.607 on 377 degrees of freedom
Multiple R-squared: 0.08109, Adjusted R-squared: 0.07621
F-statistic: 16.63 on 2 and 377 DF, p-value: $1.194 \mathrm{e}-07$

## Our first guess might be "everything is additive"



This asserts "parallel lines" for males and females

## Agency effect depends on gender?

One might imagine that rather than an additive effect of gender, as in

$$
\operatorname{posAffect}_{i}=\beta_{0}+\beta_{1} \text { agency }_{i}+\beta_{2} \text { female }_{i}
$$

it is more likely that the effect of agency differs between males and females
posAffect $_{i}=\beta_{0}+\beta_{1}$ agency $_{i}+\beta_{2}$ female $_{i}+\beta_{3}$ agency $_{i} \times$ female $^{\prime}$

```
reg.mod.6 <- lm(posAffect ~ agency*gender, data =
    dat)
```

The results are

```
summary(reg.mod.6)
```


## Agency effect depends on gender?

```
Call:
lm(formula = posAffect ~ agency * gender, data = dat)
Residuals:
    Min 1Q Median 3Q Max
-2.10608 -0.40401 0.02606 0.44460 1.32508
Coefficients:
\begin{tabular}{lrrrr} 
& Estimate & Std. Error & t value & Pr ( \(>|\mathrm{t}|\) ) \\
(Intercept) & 1.88976 & 0.22056 & 8.568 & \(2.75 \mathrm{e}-16\) \\
agency & 0.45182 & 0.08624 & 5.239 & \(2.69 \mathrm{e}-07\) \\
genderfemale & 0.62379 & 0.31834 & 1.960 & 0.0508 \\
agency: genderfemale & -0.21481 & 0.12428 & -1.728 & 0.0847
\end{tabular}
(Intercept)
agency
genderfemale
agency:genderfemale
Signif. codes:
0 '***' 0.001 ,**' 0.01 '*' 0.05 ,.' 0.1 , , 1
Residual standard error: 0.6054 on 376 degrees of freedom
Multiple R-squared: 0.08833, Adjusted R-squared: 0.08106
F-statistic: 12.14 on 3 and 376 DF, p-value: 1.331e-07
```


## Agency effect depends on gender?

## Visualize the interaction




## How to best plot that?

- My tendency has been to draw several groups on one plot.
- Others prefer "trellis" graphics, which make smaller pictures, one for each group
- In the base of R, the lattice package is provided for this purpose
- Hadley Wickham's ggplot2 package is a little bit easier to use, so we will test that.


## ggplot thumbnail sketch

- ggplot is similar in many ways to concept of R base graphics,
- We can
- draw a "blank" figure
- add pieces to it
- However,
- it uses an entirely different vocabulary, such as "geom" and "aes".
- additional graph commands do not just "draw" pieces can fundamentally alter the display.
- variable names need not be quoted (I find this confusing)


## ggplot thumbnail sketch

- The plot is initiated by a call to $\operatorname{ggplot}()$, which must specify an "aesthetic", the fundamental nature of the plot
- An interesting difference with base graphics is that we think of "adding" graph elements

```
p1 <- ggplot(data.frame, aes(...))
p1 <- p1 + new features here
p1
```

- The last p1 causes the result to be drawn in a graphic window.
- People often write a string of added-together features, but I usually test the new features one at a time.

```
p1 <- ggplot(data.frame, aes(...))
    + new feature here
    + more features
p1
```


## ggplot thumbnail sketch ...

- Sometimes the ordering of new features will make a little difference in the final display.


## ggplot blank page



```
library(ggplot2)
p1 <- ggplot(dat, aes(x = agency, y = posAffect))
p1
```


## geom_point is for points in a scatter



$$
\begin{aligned}
& \text { p1 <- p1 + geom_point(shape }=1, \text { alpha }=0.5) \\
& \text { p1 }
\end{aligned}
$$

## facet_grid() divides plot into sections



```
p1 <- p1 + geom_point(shape = 1, alpha = 0.5)
p1 <- p1 + facet_grid(. ~ gender)
p1
```


## geom_line will get line data from plotSlopes object


ps31\$newdata\$posAffect <- ps31\$newdata\$fit p1 <- p1 + geom_line (data = ps31\$newdata, color = "blue")
p1

## geom_ribbon() can draw the confidence intervals



$$
\begin{aligned}
& \text { p1 <- p1 + geom_ribbon(data = ps31\$newdata, } \\
& \text { aes(ymin = lwr, ymax = upr), } \\
& \text { fill = "pink", alpha = 0.5) } \\
& \text { p1 }
\end{aligned}
$$

## I don't want gray boxes in background!



```
p1 <- p1 + theme_bw()
p1
```


## More beautful group labels



```
p1 <- p1 + theme(strip.background =
    element_rect(color="darkgoldenrod4",
    fill="lightgoldenrod"))
p1
```


## Include ethnicity

- Previous seems to indicate there is not a "statistically significant" difference between males and females, so instead we consider ethnicity

```
reg.mod.7 <- lm(posAffect ~ agency*ethnicity +
    gender, data = dat)
summary(reg.mod.7)
```

```
Call:
lm(formula = posAffect ~ agency * ethnicity + gender, data = dat)
Residuals:
    Min 1Q Median
3Q
Max
\(-2.11727-0.38900 \quad 0.04804 \quad 0.44363 \quad 1.38301\)
Coefficients:
(Intercept)
agency
ethnicityBlack
ethnicityHispanic
ethnicityWhite
genderfemale
\begin{tabular}{rrr} 
Estimate & Std. Error & t \\
2.47790 & 0.47151 & 5.255 \\
0.15378 & 0.18164 & 0.847 \\
-0.59417 & 0.82482 & -0.720 \\
-0.20871 & 0.59406 & -0.351 \\
-0.43142 & 0.51081 & -0.845 \\
0.09997 & 0.06343 & 1.576
\end{tabular}
```


## Include ethnicity ...

```
agency:ethnicityBlack 0.29965
agency:ethnicityHispanic
0.13503
0.24453
Pr}(>|t|
(Intercept) 2.5e-07
agency 0.398
ethnicityBlack 0.472
ethnicityHispanic 0.726
ethnicityWhite 0.399
genderfemale 0.116
agency:ethnicityBlack 0.372
agency:ethnicityHispanic 0.555
agency:ethnicityWhite 0.218
---
Signif. codes:
0 '***' 0.001 '**' 0.01 '*' 0.05 ,.' 0.1 , , 1
Residual standard error: 0.6078 on 371 degrees of freedom
Multiple R-squared: 0.0933, Adjusted R-squared: 0.07375
F-statistic: 4.772 on 8 and 371 DF, p-value: 1.347e-05
```

- Again, this example is a little disappointing


## Include ethnicity ...



## ggplot trellis plot for quantile-based groups



## ggplot trellis plot for quantile-based groups ...

```
## Data must be subdivided by groups
ps71 <- plotSlopes(reg.mod.7, plotx = "agency",
    modx = "ethnicity", interval = "confidence")
ps71$newdata$posAffect <- ps71$newdata$fit
p1 <- ggplot(dat, aes(x = agency, y = posAffect))
    + geom_point(shape = 1, alpha = 0.5) +
    facet_wrap( ~ ethnicity, ncol = 2) +
        geom_line(data = ps71$newdata, color =
        "blue") +
    geom_ribbon(data = ps71$newdata, aes(ymin = lwr,
        ymax = upr), fill = "pink", alpha = 0.5) +
    theme_bw() +
    theme(strip.background =
        element_rect(color="darkgoldenrod4",
        fill="lightgoldenrod"))
p1
```


## Additive Model

```
reg.mod.10 <- lm(posAffect ~ agency + intMotiv +
    extMotiv, data = dat)
summary(reg.mod.10)
```


## Call:

lm(formula $=$ posAffect $\sim$ agency + intMotiv + extMotiv, data $=$ dat)
Residuals:

| Min | $1 Q$ | Median | 3Q | Max |
| ---: | ---: | ---: | ---: | ---: |
| -1.88002 | -0.35067 | 0.01655 | 0.42346 | 1.16862 |

Coefficients:
Estimate Std. Error $t$ value $\operatorname{Pr}(>|t|)$
(Intercept) $1.858650 .19367 \quad 9.597<2 e-16$ ***
agency $0.22583 \quad 0.06950 \quad 3.249 \quad 0.00126$ **
$\begin{array}{lrrrrr}\text { intMotiv } & 0.25207 & 0.05110 & 4.932 & 1.22 \mathrm{e}-06 & * * * \\ \text { extMotiv } & -0.07459 & 0.06629 & -1.125 & 0.26126 & \end{array}$
Signif. codes:
0 '***' 0.001 ,**' 0.01 '*' 0.05 , , 0.1 , 1

Residual standard error: 0.5877 on 376 degrees of freedom Multiple R-squared: 0.1409, Adjusted R-squared: 0.1341

## Additive Model

```
F-statistic: 20.56 on 3 and 376 DF, p-value: 2. 333e-12
```

KU

## Explore interactions

- Based on a comprehensive literature review and exhaustive theoretical analysis, the PI proposes an interaction between agency and extMotiv

```
reg.mod.11 <- lm(posAffect ~ intMotiv +
    agency*extMotiv, data = dat)
summary(reg.mod.11)
```

```
Call:
lm(formula = posAffect ~ intMotiv + agency * extMotiv, data = dat)
Residuals:
\begin{tabular}{rrrrr} 
Min & \(1 Q\) & Median & \(3 Q\) & Max \\
-1.88992 & -0.35422 & 0.01966 & 0.42660 & 1.17393
\end{tabular}
Coefficients:
\begin{tabular}{lrrrrrr} 
& Estimate & Std. Error & t value & Pr \((>|t|)\) & \\
(Intercept) & 2.75054 & 0.51547 & 5.336 & \(1.65 \mathrm{e}-07\) & \(* * *\) \\
intMotiv & 0.25260 & 0.05094 & 4.959 & \(1.08 \mathrm{e}-06\) & \(* * *\) \\
agency & -0.11041 & 0.19305 & -0.572 & 0.5677 & \\
extMotiv & -0.65218 & 0.31650 & -2.061 & 0.0400 & \(*\) \\
agency: extMotiv & 0.21506 & 0.11525 & 1.866 & 0.0628 &.
\end{tabular}
```


## Explore interactions

```
Signif. codes:
0 '***' 0.001 '***' 0.01 '*' 0.05 '.' 0.1 ' ' 1
Residual standard error: 0.5858 on 375 degrees of freedom
Multiple R-squared: 0.1488, Adjusted R-squared: 0.1398
F-statistic: 16.39 on 4 and 375 DF, p-value: 2.175e-12
```


## Explore interactions ...

- Visualize that by choosing center points of the 4 quantiles of extMotiv

$$
\begin{aligned}
& \text { ps80 <- plotSlopes(reg.mod.11, plotx = "agency", } \\
& \text { modx }=\text { "extMotiv", modxVals }=c(1.14,1.4, \\
& 1.75,3.6))
\end{aligned}
$$

## Explore interactions



## Follow-Up 1: The J-N Analysis

- When you looked at this

did you wonder the following:
(1) It looks like the black line's slope is not different from 0 , but the blue line slope certainly is.
(2) That means the "statistical significance of agency depends on the value of extMotiv."


## Follow-Up 1: The J-N Analysis ...

- Instead of asking "is agency significant?", an interaction modeler should as "are there values of extMotiv for which agency might be significant?"

That is known as a Jersey-Neyman (J-N) hypothesis analysis.
In rockchalk, find the function "testSlopes"
ps80ts <- testSlopes (ps80)

```
Values of extMotiv INSIDE this interval:
    lo hi
1.235092 20.920372
cause the slope of (b1 + b2*extMotiv) agency to be statistically
    significant
```


## Follow-Up 1: The J-N Analysis ...



## Followup 2: Nested Model Hypo Test

- A competing research team insists that we need to check interactions with intMotiv as well. This includes all interaction terms

```
reg.mod.12 <- lm(posAffect ~ agency * intMotiv *
    extMotiv, data = dat)
summary(reg.mod.12)
```

```
Call:
lm(formula = posAffect ~ agency * intMotiv * extMotiv, data = dat)
Residuals:
    Min 1Q Median
    3Q
        Max
\begin{tabular}{rrrrr} 
Min & \(1 Q\) & Median & 3Q & Max \\
-1.88373 & -0.36143 & 0.03298 & 0.40975 & 1.14755
\end{tabular}
Coefficients:
(Intercept)
\begin{tabular}{rrr} 
Estimate & Std. Error & t value \\
0.2944 & 2.1812 & 0.135 \\
1.0015 & 0.8970 & 1.117 \\
0.9717 & 0.6850 & 1.419 \\
1.1748 & 1.4165 & 0.829 \\
-0.3254 & 0.2685 & -1.212 \\
-0.5861 & 0.5547 & -1.057
\end{tabular}
```


## Followup 2: Nested Model Hypo Test ...

```
intMotiv:extMotiv -0.5460 0.4473 -1.221
agency:intMotiv:extMotiv 0.2386 0.1680 1.420
nterc
agency 0.265
intMotiv 0.157
extMotiv 0.407
agency:intMotiv 0.226
agency:extMotiv 0.291
intMotiv:extMotiv 0.223
agency:intMotiv:extMotiv 0.156
Residual standard error: 0.5861 on 372 degrees of freedom
Multiple R-squared: 0.1548, Adjusted R-squared: 0.1389
F-statistic: 9.735 on 7 and 372 DF, p-value: 3.806e-11
```

- The research question is this: Is the model with more predictors better?
- These are NESTED models (the smaller one is a simplification of the larger one).


## Followup 2: Nested Model Hypo Test ...

- A classical approach to test that is an F test, which examines the change in the sum-of-squares between the two models. The R team has bundled together a number of tests of that sort in the anova() function.

```
anova(reg.mod.10, reg.mod.11, reg.mod.12, test =
    "F")
```

```
Analysis of Variance Table
Model 1: posAffect ~ agency + intMotiv + extMotiv
Model 2: posAffect ~ intMotiv + agency * extMotiv
Model 3: posAffect ~ agency * intMotiv * extMotiv
    Res.Df RSS Df Sum of Sq F Pr (>F)
1 376 129.87
2 375 128.67 1 1.19479 3.4787 0.06295
3 372 127.77 3 0.90459}00.8779 0.45261 
```



- The comparison of models 1 and 2 is statistically significant, meaning we should keep the additional coefficients in the model


## Followup 2: Nested Model Hypo Test ...

- The comparison of models 2 and 3 is not. So the enemy research team was wrong.


## Centering and Standardizing

- We notice that many in psychology enjoy "standardized regression" or "mean-centered" regressions.
- Can do "manually", but I do that so often while teaching I created shortcuts.
- Do you want a model in which all numeric variables are centered at their means? The meanCenter function defaults to only change variables involved in interactions

```
## The mean-centered model sets the predictors at
    (x - xmean)
reg.mod.14 <- meanCenter(reg.mod.11)
summary(reg.mod.14)
```


## Centering and Standardizing ...

```
These variables were mean-centered before any transformations were made
    on the design matrix.
[1] "agencyc" "extMotivc"
The centers and scale factors were
            agencyc extMotivc
mean 2.511814 1.644807
scale 1.000000 1.000000
The summary statistics of the variables in the design matrix (after
    centering).
posAffect 3.065193 0.6315674
intMotiv
    3.022962 0.6607460
agencyc
    0.000000 0.5008115
extMotivc
    0.000000 0.4760930
agencyc:extMotivc 0.058399 0.2740184
The following results were produced from:
meanCenter.default(model = reg.mod.11)
Call:
lm(formula = posAffect ~ intMotiv + agencyc * extMotivc, data = stddat)
Residuals:
    Min 1Q Median 3Q Max
-1.88992 -0.35422 0.01966 0.42660 1.17393
```


## Centering and Standardizing ...

Coefficients:

|  | Estimate | Std. Error | t value | Pr $(>\|t\|)$ |  |
| :--- | ---: | ---: | ---: | ---: | ---: |
| (Intercept) | 2.28902 | 0.15707 | 14.574 | $<2 \mathrm{e}-16$ | $* * *$ |
| intMotiv | 0.25260 | 0.05094 | 4.959 | $1.08 \mathrm{e}-06$ | $* * *$ |
| agencyc | 0.24333 | 0.06990 | 3.481 | 0.000559 | $* * *$ |
| extMotivc | -0.11198 | 0.06905 | -1.622 | 0.105687 |  |
| agencyc:extMotivc | 0.21506 | 0.11525 | 1.866 | 0.062818 | . |

Signif. codes:
$0^{\prime * * *} 0.001^{\prime * *} 0.01^{\prime *} 0.05$,.' 0.1 , 1

Residual standard error: 0.5858 on 375 degrees of freedom Multiple R-squared: 0.1488, Adjusted R-squared: 0.1398 F-statistic: 16.39 on 4 and 375 DF, p-value: $2.175 e-12$

## - Change centerOnlyInteractors $=$ FALSE

```
reg.mod.14b <- meanCenter (reg.mod.11,
    centerOnlyInteractors = FALSE)
summary (reg.mod.14b)
```


## Centering and Standardizing ...

```
These variables were mean-centered before any transformations were made
    on the design matrix.
[1] "intMotivc" "agencyc" "extMotivc"
The centers and scale factors were
    intMotivc agencyc extMotivc
mean 3.022962 2.511814 1.644807
scale 1.000000 1.000000 1.000000
The summary statistics of the variables in the design matrix (after
    centering).
3.065193 0.6315674
intMotivc 0.000000 0.6607460
agencyc 0.000000 0.5008115
extMotivc 0.000000 0.4760930
agencyc:extMotivc 0.058399 0.2740184
The following results were produced from:
meanCenter.default(model = reg.mod.11, centerOnlyInteractors = FALSE)
Call:
lm(formula = posAffect ~ intMotivc + agencyc * extMotivc, data = stddat)
Residuals:
    Min 1Q Median 3Q Max
-1.88992 -0.35422 0.01966 0.42660 1.17393
```


## Centering and Standardizing ...

```
Coefficients:
(Intercept) 3.05263 0.03079 99.131 < 2e-16 ***
intMotivc 0.25260 0.05094 4.959 1.08e-06 ***
agencyc 0.24333 0.06990 3.481 0.000559 ***
extMotivc -0.11198 0.06905 -1.622 0.105687
agencyc:extMotivc 0.21506 0.11525 1.866 0.062818.
Signif. codes:
0 '***, 0.001 ,**' 0.01 '*'0.05 ,., 0.1 , , 1
Residual standard error: 0.5858 on 375 degrees of freedom
Multiple R-squared: 0.1488, Adjusted R-squared: 0.1398
F-statistic: 16.39 on 4 and 375 DF, p-value: 2.175e-12
```

- The coefficients hop about because of the transformation, but don't let anybody fool you. The mean-centered regression is absolutely identical to the un-centered one. Note the predicted values are identical


## Centering and Standardizing ...



## Centering and Standardizing ...

```
plot(fitted(reg.mod.14), fitted(reg.mod.11), xlab
    = "Uncentered predictors", ylab =
    "Mean-Centered predictors", main = "Predicted
    Values")
```

- The results have "seemed" different to some authors.


## Outline

## (.) Package Check!

D Check the Data

- read.table plus
- Recodes
- 


## One-Predictor Linear Regression

- The Im() function and R formula
- Access Points
- About Formulas
- Diagnostics
- The Predicted Value Framework
- Categorical Predictors
- Bug-Shooting

Add More Predictors

- Formulas
- Moderator $=$ categorical interaction
- Multi-Category factor


## Modern Applied Statistics

- The famous book by Wm. Venables and Brian Ripley, Modern Applied Statistics with S, advances the theme that statistical analysis has entered a new phase characterized the idea that
- We "interact" with estimated objects (rather than just printing output about them)
- These notes focus on linear regression modeling
- SPSS \& SAS users should notice the difference, because R makes it possible to "see inside" output objects and interrogate them in various ways


## Other regression functions

- R base also includes
- glm: generalized linear models (logit, probit, poisson, gamma)
- Recommended packages include additional regression functions
- MASS: negative binomial, Box-Cox transformation
- mgcv: generalized additive models (smoothing functions for "wiggly" model fits)
- rpart: partitioned regression trees
- Other contributed packages add many models, many of which are written in a similar style.


## References

R Core Team (2017). R: A Language and Environment for Statistical Computing. R Foundation for Statistical Computing, Vienna, Austria.

## Session

## sessionInfo ()

```
R version 3.6.0 (2019-04-26)
Platform: x86_64-pc-linux-gnu (64-bit)
Running under: Ubuntu 19.04
```

Matrix products: default
BLAS: /usr/lib/x86_64-linux-gnu/atlas/libblas.so.3.10.3
LAPACK: /usr/lib/x86_64-linux-gnu/atlas/liblapack.so.3.10.3
locale:
[1] LC_CTYPE=en_US.UTF-8 LC_NUMERIC = C
[3] LC_TIME=en_US.UTF-8 LC_COLLATE=en_US.UTF-8
[5] LC_MONETARY=en_US.UTF-8 LC_MESSAGES = en_US.UTF-8
[7] LC_PAPER=en_US.UTF-8 LC_NAME=C
[9] LC_ADDRESS=C LC_TELEPHONE=C
[11] LC_MEASUREMENT=en_US.UTF-8 LC_IDENTIFICATION=C
attached base packages:
[1] stats graphics grDevices utils datasets methods base
other attached packages:
[1] ggplot2_3.2.0 car_3.0-2 carData_3.0-2
rockchalk_1.8.144

## Session

loaded via a namespace (and not attached):
[1] lavaan_0.6-3
tidyselect_0.2.5
purrr_0.3.2
reshape2_1.4.3
[5] splines_3.6.0 colorspace_1.4-1
[9] stats4_3.6.0
[13] foreign_0.8-71
[17] plyr_1.8.4
[21] cellranger_1.1.0
[25] rio_0.5.16
[29] xtable_1.8-4
[33] mnormt_1.5-5 openxlsx_4.1.0
[37] dplyr_0.8.3 grid_3.6.0
[41] lazyeval_0.2.2 pbivnorm_0.6.0
[45] pkgconfig_2.0.2 data.table_1.12.2
[49] assertthat_0.2.1
[53] nlme_3.1-140
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rlang_0.4.0
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stringr_1.4.0
zip_2.0.2
forcats_0.4.0
scales_1.0.0
hms_0.4.2
tibble_2.1.3

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labeling_0.3
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magrittr_1.5
crayon_1.3.4

Matrix_1.2-17

R6_2.4.0
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abind_1.4-5
stringi_1.4.3
tools_3.6.0

